Logistic Regression

ACTL3142 & ACTL5110 Statistical Machine Learning for Risk Applications

Some of the figures in this presentation are taken from "An Introduction to Statistical Learning, with applications in R" (Springer, 2013) with permission from the authors: G. James, D. Witten, T. Hastie and R. Tibshirani





Lecture Outline

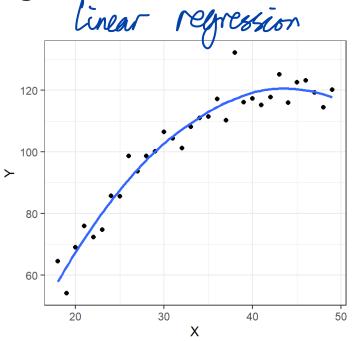
- An overview of classification
- Logistic regression Classification type problems
 Poisson regression Counting type problems
- Generalised linear models



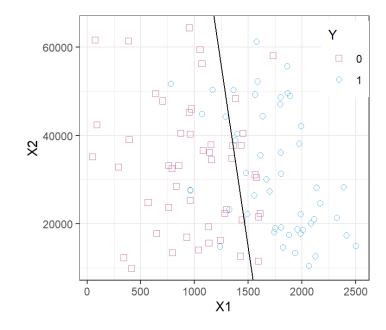


Regression vs. classification

Regression



Classification



- *Y* is quantitative continuous
- Examples: Sales prediction, claim size
 prediction, stock price modelling
- *Y* is qualitative, discrete
- Examples: Fraud detection, face recognition, accident occurrence, death





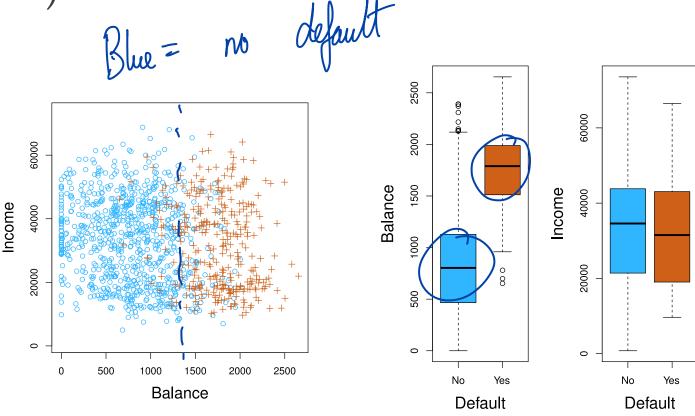
Some examples of classification problems

- Success/failure of a treatment, explained by dosage of medicine administered, patient's age, sex, weight and severity of condition, etc.
- Vote for/against political party, explained by age, gender, education level, region, ethnicity, geographical location, etc.
- Customer churns/stays depending on usage pattern, complaints, social demographics, etc.





Example: Predicting defaults (Default from ISLR2)



- default (Y) is a binary variable (yes/no or 0/1)
- Annual income (X_1) and credit card balance (X_2) may be continuous predictors



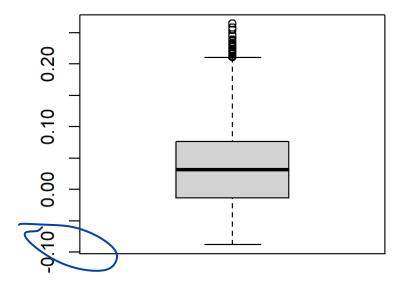


Example: Predicting defaults - Discussion

Simple linear regression on Default data:

prob- of default.

Fitted values of default probability



have regative value Non-sersical

What do you observe?





Classification problems

• Coding in the binary case is simple

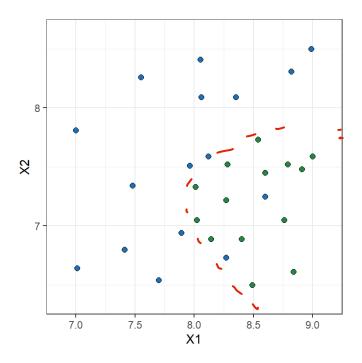
$$Y \in \{0,1\} \Leftrightarrow Y \in \{ullet,ullet\}$$

- Our objective is to find a good predictive model *f* that can:
 - 1. Estimate the probability

$$\mathbb{P}(Y=1|X)\in\{0,1\}$$

2. Classify observation

$$f(X) o \hat{Y} \in \{ullet,ullet\}$$





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Logistic regression

Extend linear regression to model binary categorical variables

$$\ln\left(rac{\mathbb{P}(Y=1|X)}{1-\mathbb{P}(Y=1|X)}
ight) = \underbrace{eta_0 + eta_1 X_1 + \cdots + eta_p X_p}_{ ext{linear model}}$$

log-odds are linear in X.

Some os Mir structure

 $log(\frac{p}{1-p})$





- Principles of Logistic Regression \mathcal{O}_{1} .

 The output is binary $Y \in \{1,0\}$
- Each case's *Y* variable has a probability between 0 and 1 that depends on the values of the predictors *X* such that

$$\mathbb{P}(Y=1|X) + \mathbb{P}(Y=0|X) = 1$$

Probability can be restated as odds

$$\mathrm{Odds}(Y=1|X) = rac{\mathbb{P}(Y=1|X)}{\mathbb{P}(Y=0|X)} = rac{\mathbb{P}(Y=1|X)}{1-\mathbb{P}(Y=1|X)}$$

• Odds are a measure of relative probabilities



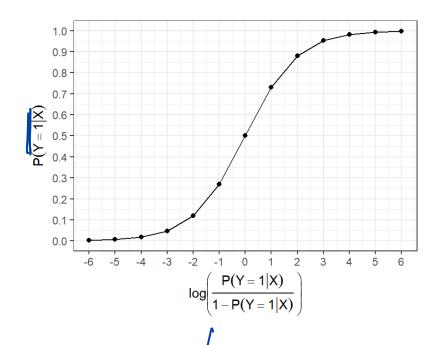


Probabilities, odds and log-odds

Goal: Transform a number between 0 and 1 into a number between $-\infty$ and $-\infty$

probability	odds	logodds
0.001	0.001	-6.907
0.250	0.333	-1.099
→ 0.500	1.000	0.000
0.750	3.000	1.099
0.999	999.000	6.907

$$P(Y=11X)=0.5$$
 $P(Y=01X)=0.5$



transformation is unique





Logistic regression

• Perform regression on log-odds

$$\ln\left(rac{\mathbb{P}(Y=1|X)}{1-\mathbb{P}(Y=1|X)}
ight)=eta_0+eta_1X_1+\cdots+eta_pX_p$$

- Use (training) data and maximum-likelihood estimation to produce estimates $\hat{\beta}_0, \hat{\beta}_1, \dots \hat{\beta}_p$.
- Predict probabilities using

$$\mathbb{P}(Y=1|X) = \frac{\mathrm{e}^{\hat{\beta}_0 + \hat{\beta}_1 X_1 + \dots + \hat{\beta}_p X_p}}{1 + \mathrm{e}^{\hat{\beta}_0 + \hat{\beta}_1 X_1 + \dots + \hat{\beta}_p X_p}}$$

$$\text{Do you know where this comes from?}$$

$$\mathbb{M}\left(\frac{\beta}{1-\beta}\right) = \mathbb{X}\beta$$





Interpretation of coefficients $\ln\left(\frac{\rho}{1-\rho}\right) = \gamma \beta - \frac{\rho}{\rho}$

$$ln(\frac{p}{1-p}) = XB - log-odds$$

• Recall for multiple linear regression we model the response as

$$Y = \beta_0 + \beta_1 X_1 + \cdots + \beta_p X_p + \varepsilon.$$

An increase of the entry $\underline{x_{ij}}$ by 1 in X we would predict Y_i to increase by $\hat{\beta}_j$ on average since

$$\mathbb{E}[Y_i|X] = \hat{eta}_0 + \hat{eta}_1 x_{i1} + \cdots + \hat{eta}_j (x_{ij}+1) + \cdots + \hat{eta}_p x_{ip}$$

- For **logistic regression** we have a similar relationship. When x_{ij} increases by 1 we would expect the **log-odds** for Y_i to increase by β_i .
- The new predicted probability of success by increasing x_{ij} by 1 is now

$$\mathbb{P}(Y_i = 1|X) = rac{\mathrm{e}^{\hat{eta}_0 + \hat{eta}_1 x_{i1} + \cdots + \hat{eta}_j (x_{ij} + 1) + \cdots + \hat{eta}_p x_{ip}}}{1 + \mathrm{e}^{\hat{eta}_0 + \hat{eta}_1 x_{i1} + \cdots + \hat{eta}_i (x_{ij} + 1) + \cdots + \hat{eta}_p x_{ip}}}.$$

Convince yourself that the probability does increase if β_i is positive!





How are the coefficients estimated?

• Recall the Bernoulli distribution is parameterised by a parameter p and has the density

$$f(y) = p^y (1-p)^{1-y}.$$

• In logistic regression we maximise the likelihood of the data. Denote

$$p(y_i; eta)
eq \overbrace{1 - \mathrm{e}^{-\mathrm{x}_i eta}}^1,$$

where x_i denotes the i'th row of X.

• We maximise the log-likelihood below

$$\ell(eta) = \sum_{i=1}^n y_i \ln p(y_i;eta) + (1-y_i) \ln (1-p(y_i;eta)).$$

We take partials w.r.t. to each β_j and set to 0. Needs numerical approximation.

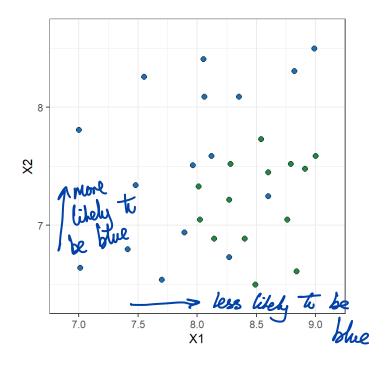




Toy example: Logistic Regression

$$Y = egin{cases} rac{1 & ext{if} ullet}{0 & ext{if} ullet} & & \ln\left(rac{\mathbb{P}(Y=1|X)}{1-\mathbb{P}(Y=1|X)}
ight) = eta_0 + eta_1 X_1 + eta_2 X_2 \end{cases}$$

- The parameter estimates are $\hat{\beta}_0 = 13.671$, $\hat{\beta}_1 = -4.136$, $\hat{\beta}_2 = 2.803$
- $\hat{\beta}_1 = -4.136$ implies that the bigger X_1 the lower the chance it is a blue point
- $\hat{\beta}_2 = 2.803$ implies that the bigger X_2 the higher the chance it is a blue point



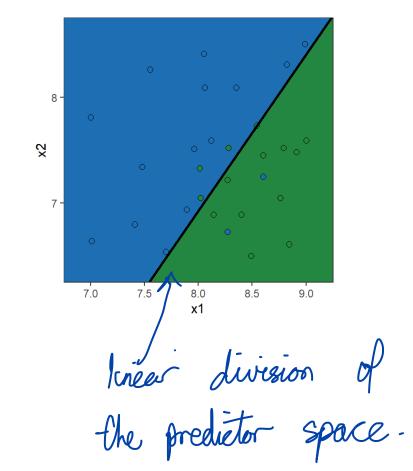




Toy example: Logistic Regression

$$\ln\left(\frac{\mathbb{P}(Y=1|X)}{1-\mathbb{P}(Y=1|X)}\right) = 13.671-4.136X_1+2.803X_2$$
 log odds

X 1	X2	log-odds	$P(Y=1 \mid X)$	prediction
7.0	8.0	7.14	0.9992	blue
8.0	7.5	1.61	0.8328	blue
8.0	7.0	0.20	0.5508	blue
8.5	7.5	-0.46	0.3864	green
9.0	7.0	-3.93	0.0192	green
)	







Some important points about logistic regression

- Changes in predictor values correspond to changes in the *log-odds*, not the probability
- Evaluating predictors to add / remove is the same as in <u>linear regression</u>. The only change is the form of the response
- As a result, most of the modelling limitations of linear regression (e.g. collinearity) carry over as well
- Possible to do logistic regression on non-binary responses, but not used that often, and not covered here





Example: Predicting defaults

```
glmStudent <- glm default ~ student, family = binomial(), data = ISLR2::Default)</pre>
        2 summary(glmStudent)
    Call:
    glm(formula = default ~ student, family = binomial(), data = ISLR2::Default)
    Coefficients:
    Estimate Std. Error z value Pr(>|z|)
(Intercept) -3.50413    0.07071   -49.55 < 2e-16 ***
studentYes    0.40489    0.11502    3.52    0.000431 ***
    Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
     (Dispersion parameter for binomial family taken to be 1)
    Null deviance: 2920.6 on 9999 degrees of freedom
Residual deviance: 2908.7 on 9998 degrees of freedom

Week
    AIC: 2912.7
    Number of Fisher Scoring iterations: 6
If default represents Y=1, you are more likely to default if you are a students.
```



Example: Predicting defaults

```
glmAll <- glm(default ~ balance + income + student, family = binomial(), data = ISLR2::Default)</pre>
   2 summary(glmAll)
Call:
glm(formula = default ~ balance + income + student, family = binomial(),
   data = ISLR2::Default)
Coefficients:
            Estimate Std. Error z value Pr(>|z|)
(Intercept) -1.087e+01 4.923e-01 -22.080 < 2e-16 ***
           5.737e-03 2.319e-04 24.738 < 2e-16 ***
balance
           3.033e-06 8.203e-06 0.370 0.71152
income
studentYes -6.468e-01 2.363e-01 -2.738 0.00619 **
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for binomial family taken to be 1)
   Null deviance: 2920.6 on 9999 degrees of freedom
Residual deviance: 1571.5 on 9996 degrees of freedom
AIC: 1579.5
Number of Fisher Scoring iterations: 8
  In the presence of balance and income, you are a student.
```





Example: Predicting defaults - Discussion

Results of logistic regression:

default against student

Predictor	Coefficient	Std error Z-statist		P-value
(Intercept)	-3.5041	0.0707	-49.55	<0.0001
student = Yes	0.4049	0.1150	3.52	0.0004

default against balance, income, and student

Predictor	Coefficient	Std error	Z -statistic	P-value
(Intercept)	-10.8690	0.4923	-22.080	< 0.0001
balance	0.0057	2.319e-04	24.738	< 0.0001
income	0.0030	8.203e-06	0.370	0.71152
student = Yes	-0.6468	0.2362	-2.738	0.00619





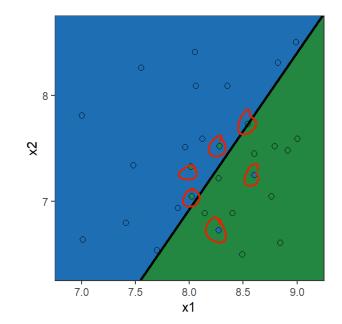
Assessing accuracy in classification problems

• We assess model accuracy using the error rate

error rate
$$=\frac{1}{n}\sum_{i=1}^{n}\underbrace{I(y_{i}\neq\hat{y}_{i})}_{i=1}$$

• In our toy example with a 50% threshold

training error rate
$$=\frac{6}{30}=0.2$$





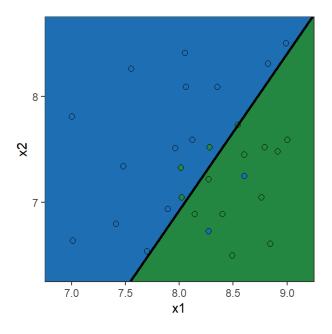


Confusion matrix: Toy example (50% Threshold)

• Confusion matrix

	Y = 0	Y = 1	Total
$\hat{Y} = 0$	10	2	12
$\hat{Y}=1$	4	14	18
Total	14	16	30

- True-Positive Rate = $\frac{14}{16}$ = 0.875
- False-Positive Rate = $\frac{4}{14} = 0.286$







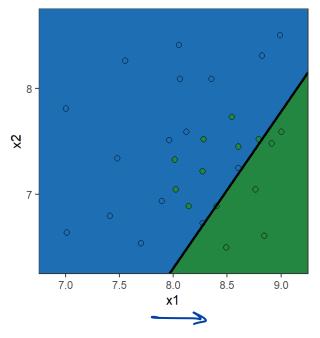
Confusion matrix: Toy example (15% Threshold)

• Confusion matrix

	Y = 0	Y = 1	Total
$\hat{Y} = 0$	6	0	6
$\hat{Y} = 1$	8	16	24
Total	14	16	30

- True-Positive Rate = $\frac{16}{16} = 1$
- False-Positive Rate = $\frac{8}{14} = 0.429$

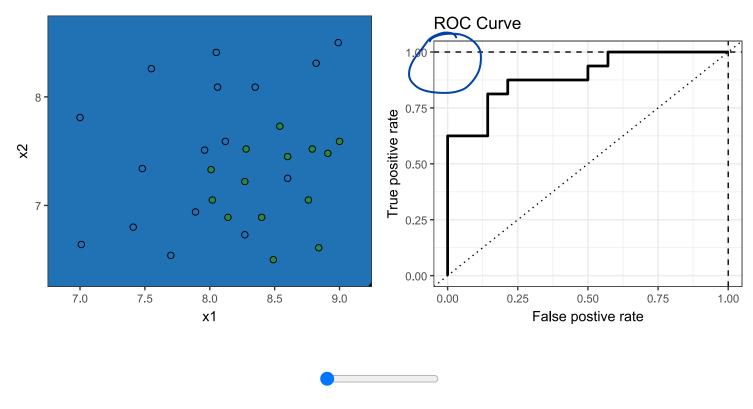
larger false positive rate







ROC Curve and AUC: Toy example



- ROC Curve: Plots the true-positive rate against the false-positive rate
- A good model will have its ROC curve hug the top-left corner more
- AUC is the area under the ROC curve: For this toy example AUC= 0.8929





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- An overview of classification
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- Poisson regression
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Poisson regression - Motivation

In many application we need to model <u>count data</u>:

- In mortality studies the aim is to explain the number of deaths in terms of variables such as age, gender and lifestyle.
- In health insurance, we may wish to explain the number of claims made by different individuals or groups of individuals in terms of explanatory variables such as age, gender and occupation.
- In general insurance, the count of interest may be the number of claims made on vehicle insurance policies. This could be a function of the color of the car, engine capacity, previous claims experience, and so on.





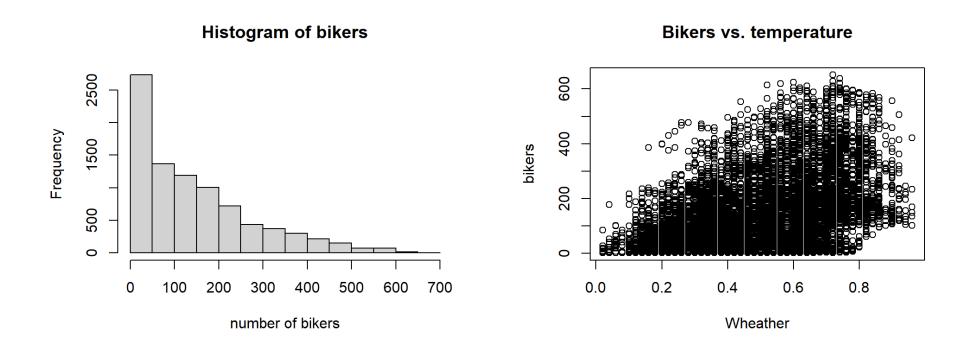
The Bikeshare dataset

```
1 str(ISLR2::Bikeshare)
'data.frame': 8645 obs. of 15 variables:
$ season
         : num 111111111...
          : Factor w/ 12 levels "Jan", "Feb", "March", ...: 1 1 1 1 1 1 1 1 1 1 ...
$ mnth
$ dav
          : num 111111111...
          : Factor w/ 24 levels "0","1","2","3",..: 1 2 3 4 5 6 7 8 9 10 ...
$ hr
$ holiday : num 0000000000...
$ weekday : num 6666666666...
$ workingday: num 0000000000...
$ weathersit: Factor w/ 4 levels "clear", "cloudy/misty", ..: 1 1 1 1 1 2 1 1 1 1 ...
           : num 0.24 0.22 0.22 0.24 0.24 0.24 0.22 0.2 0.24 0.32 ...
$ temp
           : num 0.288 0.273 0.273 0.288 0.288 ...
$ atemp
           : num 0.81 0.8 0.8 0.75 0.75 0.75 0.8 0.86 0.75 0.76 ...
$ hum
$ windspeed : num  0  0  0  0  0  0.0896  0  0  0  0  ...
$ casual
          : num 3 8 5 3 0 0 2 1 1 8 ...
$ registered: num 13 32 27 10 1 1 0 2 7 6 ...
$ bikers
          : num 16 40 32 13 1 1 2 3 8 14 ...
```





The Bikeshare dataset - Discussion



How could we model the number of bikers as function of the other variables?





Why not use muliple linear regression?

$$Y = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p + \epsilon$$

- Could predict negative values
- Constant variance may be inadequate
- Assumes continuous numbers while counts are integers

$$\log(Y) = eta_0 + eta_1 X_1 + \dots + eta_p X_p + \epsilon$$

- Solves problem of negative values
- May solve constant variance problem
- Assumes continuous numbers while counts are integers
- Not applicable with zero counts





Poisson regression

• Assume that $Y \sim \text{Poisson}(\underline{\lambda})$

$$\mathbb{P}(Y=k) = rac{e^{\lambda}\lambda^k}{k!} \quad ext{for } k=0,1,2,\dots \quad ext{with } \mathbb{E}[Y] = ext{Var}(Y) = \lambda$$

• Assume that $\mathbb{E}[Y] = \lambda(X_1, \dots, X_p)$ is log-linear in the predictors

$$\log(\lambda(X_1,\ldots,X_p))=eta_0+eta_1X_1+\cdots+eta_pX_p$$

• Use data and maximum-likelihood estimation to obtain $\hat{\beta}_0, \hat{\beta}_1, \dots \hat{\beta}_p$

$$\mathcal{L}(eta_0,eta_1,\ldots,eta_p) = \prod_{i=1}^n rac{e^{\lambda(x_i)}\lambda(x_i)^{y_i}}{y_i!} \quad ext{with} \quad \lambda(x_i) = eta_0 + eta_1x_{i1} + \cdots + eta_px_{p1}.$$





Some important points about Poisson regression



- Interpretation: An increase in X_j by one unit is associated with a change in $\mathbb{E}[Y]$ by a factor e^{β_j} .
- Mean-variance relationship: $\mathbb{E}[Y] = \text{Var}(Y) = \lambda$ implies that the variance is non-constant and increases with the mean.
- Non-negative fitted values: Predictions are always positive
- Evaluating predictors to add / remove is the same as in linear regression. The only change is the form of the response
- As a result, most of the modelling limitations of linear regression (e.g. collinearity) carry over as well

$$log(x) = \beta_0 + \beta_1 x_1 + \cdots + \beta_p x_p.$$





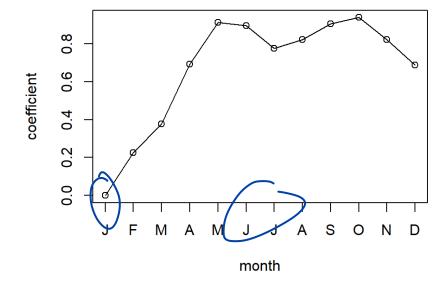
Poisson regression - Bikeshare dataset

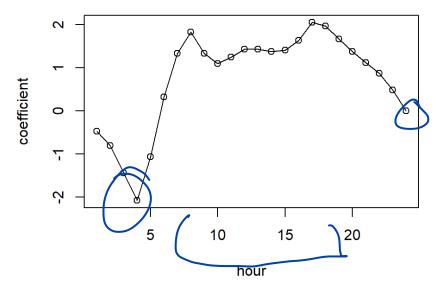
```
glm(formula = bikers ~ workingday + temp + weathersit + mnth +
    hr, family = poisson(), data = ISLR2::Bikeshare)
Coefficients:
                           Estimate Std. Error z value Pr(>|z|)
 (Intercept)
                                      0.009720 277.124 < 2e-16 ***
                           2.693688
workingday
                           0.014665
                                      0.001955
                                                 7.502 6.27e-14 ***
                                                 68.434 < 2e-16 ***
                           0.785292
                                      0.011475
temp
                          -0.075231
weathersitcloudy/misty
                                      0.002179
                                                -34.528 < 2e-16 ***
weathersitlight rain/snow
                          -0.575800
                                      0.004058 -141.905 < 2e-16 ***
weathersitheavy rain/snow
                          -0.926287
                                      0.166782
                                                -5.554 2.79e-08 ***
mnthFeb
                           0.226046
                                      0.006951
                                                 32.521 < 2e-16 ***
mnthMarch
                           0.376437
                                      0.006691
                                                 56.263 < 2e-16 ***
mnthApril
                           0.691693
                                      0.006987
                                                 98.996 < 2e-16 ***
                           0.910641
mnthMay
                                      0.007436 122.469 < 2e-16 ***
                           0.893405
                                      0.008242 108.402 < 2e-16 ***
mnthJune
                           0.773787
                                                 87.874 < 2e-16 ***
mnthJuly
                                      0.008806
mnthAug
                           0.821341
                                      0.008332
                                                 98.573 < 2e-16 ***
                           0.903663
                                      0.007621
                                                118.578 < 2e-16 ***
mnthSept
                           0.937743
                                      0.006744 139.054 < 2e-16 ***
mnthOct
```





Poisson regression - Bikeshare dataset









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Generalised linear models

	Linear Regression	Logistic Regression	Poisson Regression	Generalised Linear Models
 Type of Data	Continuous	Binary (Categorical)	Count	Flexible
 Use	Prediction of continuous variables	Classification	Prediction of the number of events	Flexible
Distribution of Y	Normal	Bernoulli (Binomial for multiple trials)	Poisson	Exponential Family
$\mathbb{E}[Y X]$	$X\beta$	$rac{e^{Xeta}}{1+e^{Xeta}}$	e^{Xeta}	$g^{-1}(Xeta)$
Link Function Name	Identity	Logit	Log	Depends on the choice of distribution
Link Function Expression	$\eta(\mu)=\mu$	$\eta(\mu) = \log\left(rac{\mu}{1-\mu} ight)$	$\eta(\mu) = \log(\mu)$	Depends on the choice of distribution



